# On the design and analysis of branching schemes with killing for rare event Monte Carlo estimation 

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## Outline

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## Problem of Interest and LD Scaling

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Assume $\left\{X^{n}(\cdot)\right\}$ satisfies a Large Deviation Principle with rate function

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I_{T}(\phi)=\int_{0}^{T} L(\phi, \dot{\phi}) d t
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if $\phi$ is AC and $I_{T}(\phi)=\infty$ else. Heuristically, for $T<\infty$, given $\phi$, small $\delta>0, x_{n} \rightarrow x=\phi(0)$ and large $n$

$$
P_{x_{n}}\left\{\sup _{0 \leq t \leq T}\left\|X^{n}(t)-\phi(t)\right\| \leq \delta\right\} \approx e^{-n I_{T}(\phi)}
$$

## Problem of Interest and LD Scaling

To estimate:

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E_{X_{n}}\left[\sum_{i=0}^{\tau^{n}} e^{-n F\left(X_{i}^{n}\right)}\right], \text { where } \tau^{n} \doteq \inf \left\{i: X_{i}^{n} \in M\right\}
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Example:
$M=A \cup B, B$ rare, $A$ typical, and $F(x)=0, x \in B, F(x)=\infty$ otherwise.

## Problem of Interest and LD Scaling



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Under conditions of regularity on $F$ and bounds on $\tau^{n}$ :
$\begin{aligned}-\frac{1}{n} \log E_{X_{n}}\left[\sum_{i=0}^{\tau^{n}} e^{-n F\left(X_{i}^{n}\right)}\right] & \rightarrow \inf \left\{I_{T}(\phi)+F(\phi(T)): \phi(0)=x, T<\infty\right\} \\ & =\gamma .\end{aligned}$

## Some Estimation Generalities

(1) General approach: construct iid random variables $\theta_{1}^{n}, \ldots, \theta_{K}^{n}$ with $E \theta_{1}^{n}=E_{X_{n}}\left[\sum_{i=0}^{\tau^{n}} e^{-n F\left(X_{i}^{n}\right)}\right]$ and use the unbiased estimator

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\hat{q}_{n, K}\left(x_{n}\right) \doteq \frac{\theta_{1}^{n}+\cdots+\theta_{K}^{n}}{K}
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-\frac{1}{n} \log E\left(\theta_{1}^{n}\right)^{2} \leq-\frac{2}{n} \log E \theta_{1}^{n}=-\frac{2}{n} \log E_{X_{n}}\left[\sum_{i=0}^{\tau^{n}} e^{-n F\left(X_{i}^{n}\right)}\right] \rightarrow 2 \gamma
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(9) An estimator is called asymptotically efficient if

$$
\liminf _{n \rightarrow \infty}-\frac{1}{n} \log E\left(\theta_{1}^{n}\right)^{2} \geq 2 \gamma
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A single particle is started at $x$ that follows the same law as $X^{n}$, but branches into a number of independent copies each time a new level is reached.

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The number of new particles $S$ can be random (though independent of past data), and a multiplicative weight $w_{i}$ is assigned to the $i$ th descendent, where

$$
E \sum_{i=1}^{S} w_{i}=1
$$

## Splitting-Type Schemes

Evolution continues until every particle hits $M$. Let
$N_{x}^{n}=$ number of particles generated
$X_{i}^{n}(j)=$ trajectory of $j$ th particle,
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Then

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Problems

- If splitting is too infrequent, do not explore state space (standard Monte Carlo).
- If too frequent, we have exponential growth in number of surviving particles.


## RESTART and DPR

An obvious inefficiency-continuing trajectories far from the places of interest. When killing trajectories, care needed to avoid bias.

- RESTART (REpetitive Simulation Trials After Reaching Thresholds, due to Villen-Altamirano and Villen-Altamirano).


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- Particles can jump multiple thresholds $(j \rightarrow k)$.
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- Same problems as with ordinary splitting, but analysis much more difficult due to dependence on threshold of birth.


## Implementation Via Importance Functions

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and a mean increase in number of particles per threshold of

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& \left.\times\left[\sum_{j=i}^{\tau^{1, i, n}} e^{-n F\left(X_{j}^{1, i, n}\right)}\right]\left[\sum_{j=i}^{\tau^{2, i, n}} e^{-n F\left(X_{j}^{2, i, n}\right)}\right]\right]
\end{aligned}
$$

where $X_{j}^{k, i, n}$ are (conditionally) independent copies of $X_{j}^{n}$ that start at $X_{i}^{n}$ at $j=i$.

## Statement of Main Results

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Then

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- If $U$ is a subsolution

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\lim _{n \rightarrow \infty}-\frac{1}{n} \log E\left[\left(\theta^{n}\right)^{2}\right]=\inf _{y}\{\mathcal{J}(x, y)+(U(x)-U(y)) \vee 0+2 F(y)\}
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Achieved, and hence asymptotic optimality, if at minimizing $y$

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U(x)-U(y)=\mathcal{J}(x, y)
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## The F-V Quasipotential and Subsolutions

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A special case. Product form or asymptotically product form stochastic networks, $Q(y)=\langle a, y\rangle$.

## Example, Tandem Queue

$$
U(y)=-\left[\log \left(\frac{\mu_{1}}{\lambda}\right)\right] y_{1}-\left[\log \left(\frac{\mu_{2}}{\lambda}\right)\right] y_{2},
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$\lambda=1, \mu_{1}=\mu_{2}=4.5$.

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Shared buffer capacity $n$ :

| $n$ | 30 | 40 | 50 |
| :---: | :---: | :---: | :---: |
| Theoretical Value | $2.63 \times 10^{-18}$ | $1.03 \times 10^{-24}$ | $3.80 \times 10^{-31}$ |
| Estimate | $2.63 \times 10^{-18}$ | $1.06 \times 10^{-24}$ | $3.83 \times 10^{-31}$ |
| Std. Err. | $0.08 \times 10^{-18}$ | $0.04 \times 10^{-24}$ | $0.15 \times 10^{-31}$ |
| $95 \%$ C.I. | $[2.47,2.79] \times 10^{-18}$ | $[0.99,1.14] \times 10^{-24}$ | $[3.54,4.13] \times 10^{-31}$ |
| Time Taken (s) | 3 | 6 | 8 |

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Separate buffers each of capacity $n$ :

| $n$ | 10 | 20 | 30 |
| :---: | :---: | :---: | :---: |
| Theoretical Value | $9.64 \times 10^{-8}$ | $1.60 \times 10^{-15}$ | $2.64 \times 10^{-23}$ |
| Estimate | $9.70 \times 10^{-8}$ | $1.57 \times 10^{-15}$ | $2.64 \times 10^{-23}$ |
| Std. Err. | $0.16 \times 10^{-8}$ | $0.03 \times 10^{-15}$ | $0.06 \times 10^{-23}$ |
| $95 \%$ C.I. | $[9.39,10.0] \times 10^{-8}$ | $[1.51,1.63] \times 10^{-15}$ | $[2.53,2.75] \times 10^{-23}$ |
| Time Taken (s) | 3 | 12 | 26 |

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- There is an analogous theory for importance sampling, but it imposes stronger conditions on the subsolution. Differences may be significant.

