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Asymptotic distributions of some scale estimators  
in nonlinear models with long memory errors  
having infinite variance

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- ▶ An extremely important and general statistical model

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- For this, the regression residuals in the definition of M-estimator must be divided by scale estimator. See Koul (2002).
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- The fact that each of these estimators estimates a different scale parameter is not a point of concern if our goal is only to use them in arriving at scale invariant robust estimators of  $\beta_0$ .

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In addition, suppose  $\widehat{\beta}$  is an estimator of  $\beta_0$  satisfying

$$\|n^{1-d-1/\alpha}(\widehat{\beta} - \beta_0)\| = O_p(1). \quad (9)$$

## 4. Asymptotic distributions of scale estimators

Recall:  $s_1 := \text{med}\{|r_{ni}|; 1 \leq i \leq n\}$ ,  $s_2 := \text{med}\{|r_{ni} - r_{nj}|; 1 \leq i < j \leq n\}$   
where  $r_{ni} := X_{ni} - g(\widehat{\beta}, z_{ni})$  are residuals of regression model in (1)

$F(x) = P(\varepsilon_i \leq x) = \text{d.f. of errors}$ ,  $f(x) = F'(x)$ ,  $f_{\pm}(x) := f(x) \pm f(-x)$

Let

$$\alpha_* := \alpha(1 - d).$$

Note

$$1 < \alpha_* < \alpha \quad \text{for} \quad 0 < d < 1 - 1/\alpha, 1 < \alpha < 2.$$

**Thm 1** Suppose regression model (1) holds with regression function satisfying Assumption G( $a_n$ ) with  $a_n = n^{1-d-1/\alpha}$  and errors satisfying Assumption E( $\alpha, d$ ) with  $1 < \alpha < 2, 0 < d < 1 - 1/\alpha$ .

In addition, suppose  $\widehat{\beta}$  is an estimator of  $\beta_0$  satisfying

$$\|n^{1-d-1/\alpha}(\widehat{\beta} - \beta_0)\| = O_p(1). \quad (9)$$

(i) Let  $f(\sigma_1) \neq f(-\sigma_1)$ . Then, for every  $x \in \mathbb{R}$ ,

$$\begin{aligned} & P(n^{1-d-1/\alpha}(s_1 - \sigma_1) \leq x\sigma_1) \\ &= P\left(n^{1-d-1/\alpha}\left(\bar{\varepsilon}_n + \left(\frac{1}{n} \sum_{i=1}^n \dot{g}(\beta_0, z_{ni})\right)'(\hat{\beta} - \beta_0)\right) \geq -\frac{x\sigma_1 f_+(\sigma_1)}{f_-(\sigma_1)}\right) + o(1). \end{aligned}$$

(ii) Let  $f(\sigma_1) = f(-\sigma_1)$ . Then, for every  $x \in \mathbb{R}$ ,

$$P(n^{1-1/\alpha_*}(s_1 - \sigma_1) \leq x\sigma_1) \rightarrow P(Z_1^* \leq x\sigma_1 f_+(\sigma_1)),$$

where  $Z_1^* := \mathcal{Z}^*(\sigma_1) - \mathcal{Z}^*(-\sigma_1)$  and  $\mathcal{Z}^*(x), x \in \mathbb{R}$  is  $\alpha_*$ -stable process defined in (13) below.

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The last result is a consequence of the following URP for the EP:

### Thm 3 (URP I for the EP)

**Thm 3** (URP I for the EP) Suppose  $\{\varepsilon_i\}$  satisfies Assumption E( $\alpha, d$ ), for  $0 < d < 1 - 1/\alpha, 1 < \alpha < 2$ ,  $\bar{\varepsilon}_n = n^{-1} \sum_{i=1}^n \varepsilon_i$ .

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$$\psi_{\pm}(x) := (c_0^{\frac{1}{1-d}} / (1-d)) \int_0^{\infty} (F(x \mp s) - F(x) \pm f(x)s) s^{-1-\frac{1}{1-d}} ds \quad (14)$$

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**Thm 5** (URP II) Under the same conditions as in Thms 3 and 4,

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- The problem of the limit distribution of  $\sum_{i=1}^n |\varepsilon_i|^p$  for LM infinite variance moving averages  $\{\varepsilon_i\}$  is related to that of the limit distributions of power variations of semi-stationary Lévy process discussed in Basse-O'Connor, Lachièze-Rey and Podolskij (2015)

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$$\lim_{z \rightarrow \pm\infty} |z|^{-1/(1-d)} \eta(x; z) = \psi_{\pm}(x) = \text{const} \int_0^{\infty} \left( F(x \mp s) - F(x) \pm f(x)s \right) \frac{ds}{s^{1+\frac{1}{1-d}}}$$

*Step 3.* Show Step 2 and  $\alpha$ -tails of  $\zeta_s$  imply  $\alpha_*$ -tails of  $\eta(x; \zeta_s)$

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### **Sketch of proof of Thm 5 (URP II).**

*Step 1.* Decompose

$$F_n(x) - F(x) + f(x)\bar{\varepsilon}_n - \mathcal{Z}_n(x)$$



*Step 3.* Show Step 2 and  $\alpha$ -tails of  $\zeta_s$  imply  $\alpha_*$ -tails of  $\eta(x; \zeta_s)$  and hence  $\alpha_*$ -stable limit of  $\sum_{s=1}^n \eta_{n,s}(x; \zeta_s)$  for  $x$  fixed (also convergence of finite-dimensional distributions)

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$$\begin{aligned} F_n(x) - F(x) + f(x)\bar{\varepsilon}_n - \mathcal{Z}_n(x) \\ = n^{-1} \sum_{i=1}^n (I(\varepsilon_i \leq x) - F(x) + f(x)\varepsilon_i) - \mathcal{Z}_n(x) \end{aligned}$$

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$$E|W_{n,s,v}(x, y)|^r \leq \mu(x, y) \sum_{i=1 \vee s}^n |b_{i-s}|^r |b_{i-v}|^r$$

## References

- ▶ Astrauskas, A. (1983). Limit theorems for sums of linearly generated random variables. *Lithuanian Math. J.*, **23**, 127–134.
- ▶ Astrauskas, A., Levy, J.B. and Taqqu, M.S. (1991). The asymptotic dependence structure of the linear fractional Lévy motion. *Lithuanian Math. J.*, **31**, 1–28.
- ▶ Avram, F. and Taqqu, M.S. (1986). Weak convergence of moving averages with infinite variance. In: Eberlein, E. and M.S. Taqqu (eds), *Dependence in Probability and Statistics*, pp. 399-415. Birkhäuser, Boston.
- ▶ Avram, F. and Taqqu, M.S. (1992). Weak convergence of sums of moving averages in the  $\alpha$ -stable domain of attraction. *Ann. Probab.*, **20**, 483–503.
- ▶ Basse-O'Connor, A., Lachièze-Rey, R. and Podolskij, M. (2015). Limit theorems for stationary increments Lévy driven moving averages. Preprint.
- ▶ Dehling, H. and Taqqu, M.S. (1989). The empirical process of some long range dependent sequences with an application to U-statistics. *Ann. Statist.*, **17**, 1767–1783.
- ▶ Dobrushin, R.L. and Major, P. (1979) Non-central limit theorems for non-linear functionals of Gaussian fields. *Probab. Th. Rel. Fields* **50**, 27–52.
- ▶ Giraitis, L. and Surgailis, D. (1999). Central limit theorem for the empirical process of a linear sequence with long memory. *J. Statist. Plan. Inf.* **80**, 290–311.
- ▶ Giraitis, L., Koul, H.L. and Surgailis D. (1996). Asymptotic normality of regression estimators with long memory errors. *Statist. Probab. Letters* **29**, 317-335.
- ▶ Giraitis, L., Koul, H.L. and Surgailis, D. (2012). *Large Sample Inference for Long Memory Processes*. Imperial College Press, London.

- ▶ Ho, H.-C. and Hsing, T. (1996). On the asymptotic expansion of the empirical process of long memory moving averages. *Ann. Statist.*, **24**, 992-1024, 1996.
- ▶ Hsing, T. (1999) On the asymptotic distributions of partial sums of functionals of infinite-variance moving averaged. *Ann. Probab.*, **27**, 1579-1599.
- ▶ Huber, P.J. (1981). *Robust Statistics*. Wiley, New York.
- ▶ Hult, H. and Samorodnitsky, G. (2008). Tail probabilities for infinite series of regularly varying random vectors. *Bernoulli*, **14**, 838-864.
- ▶ Ibragimov, I.A. and Linnik, Yu.V. (1971). *Independent and Stationary Sequences of Random Variables*. Wolters-Noordhoff, Groningen.
- ▶ Kasahara, Y. and Maejima, M. (1988). Weighted sums of i.i.d. random variables attracted to integrals of stable processes. *Probab. Th. Rel. Fields*, **78**, 75-96.
- ▶ Kokoszka, P.S. and Taqqu, M.S. (1995). Fractional ARIMA with stable innovations. *Stoch. Proc. Appl.* **60**, 19-47.
- ▶ Koul, H.L. (2002). Asymptotic distributions of some scale estimators in nonlinear models. *Metrika*, **55**, 75-90.
- ▶ Koul, H.L. (2002a). *Weighted Empirical Processes in Dynamic Nonlinear Models. 2nd Edition. Lecture Notes Series in Statistics*, **166**, Springer, New York, N.Y., USA.
- ▶ Koul, H.L. and Surgailis, D. (2001). Asymptotics of empirical processes of long memory moving averages with infinite variance. *Stochastic Process. Appl.*, **91**, 309-336.
- ▶ Koul, H.L. and Surgailis, D. (2002). Asymptotic expansion of the empirical process of long memory moving averages. In: H. Dehling, T. Mikosch and M. Sorensen (eds.), *Empirical Process Techniques for Dependent Data*, pp. 213-239. Birkhäuser: Boston.



- ▶ Koul, H.L., Baillie, R. and Surgailis, D. (2004). Regression model fitting with a long memory covariate process. *Econometric Theory*, **20**, 485–512.
- ▶ Koul, H.L. and Surgailis, D. (2017). Asymptotic distributions of some scale estimators in nonlinear models with long memory errors having infinite variance. Preprint.
- ▶ Rosenblatt, M. (1961). Independence and dependence. *Proceed. 4th Berk. Symp. Math. Statist. & Probab.* **2** 431-443. University of California Press, Berkeley.
- ▶ Samorodnitsky, G. and Taqqu, M.S. (1994). *Stable Non-Gaussian Random Processes*. Chapman and Hall, New York.
- ▶ Surgailis, D. (2002). Stable limits of empirical processes of long memory moving averages with infinite variance. *Stochastic Process. Appl.*, **100**, 255–274.
- ▶ Surgailis, D. (2004). Stable limits of sums of bounded functions of long memory moving averages with finite variance. *Bernoulli*, **10**, 327–355.
- ▶ Taqqu, M.S. (1975). Weak convergence to Fractional Brownian Motion and to the Rosenblatt Process. *Z. Wahrsch. verw. Geb.* **31**, 287–302.
- ▶ Taqqu, M.S. (1979). Convergence of integrated processes of arbitrary Hermite rank. *Z. Wahrsch. verw. Geb.* **50**, 53-83.