

Tail probabilities of St. Petersburg sums, trimmed sums, and their limit

Péter Kevei

TU Munich

Conference on Ambit Fields and Related Topics
Aarhus, Denmark

Outline

St. Petersburg game

Sum

Maximum

Conditioning on the maximum

Number of maximum terms

Conditional limit results

Trimmed sums

Finite number of summands

Properties of the r -trimmed limit

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St. Petersburg paradox

Nicolaus Bernoulli (1713): Paul's gain X , then

$$\mathbf{P}\{X = 2^k\} = \frac{1}{2^k}, \quad k = 1, 2, \dots$$

What is the fair price?

Paradox:

$$\mathbf{E}(X) = \sum_{k=1}^{\infty} 2^k \frac{1}{2^k} = \sum_{k=1}^{\infty} 1 = \infty$$

$$\text{but } \mathbf{P}\{X > 40\} = 2^{-5} = 0.03125$$

‘there ought not be a sane man who would not happily sell his chance for forty ducats’ – Nicolaus Bernoulli

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LLN

$X_1, X_2, \dots$ iid St. Petersburg rv's $S_n = \sum_{k=1}^n X_k$

Theorem (Feller (1945))

$$\frac{S_n}{n \log_2 n} \xrightarrow{\mathbf{P}} 1$$

There are **no** strong laws!

Theorem (Adler (1990), Chow & Robbins (1961))

$$\liminf_{n \rightarrow \infty} \frac{S_n}{n \log_2 n} = 1 \text{ a.s.}, \quad \limsup_{n \rightarrow \infty} \frac{S_n}{n \log_2 n} = \infty \text{ a.s.}$$

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CLT

$$\frac{S_n - c_n}{a_n} \xrightarrow{\mathcal{D}} ?$$

Doeblin–Gnedenko criterion:

$$\mathbf{P}\{X \leq x\} = \begin{cases} 0, & \text{for } x < 2, \\ 1 - 2^{-\lfloor \log_2 x \rfloor} = 1 - \frac{2^{\{\log_2 x\}}}{x}, & \text{for } x \geq 2, \end{cases}$$

$2^{\{\log_2 x\}}$ is not slowly varying ($\{\cdot\}$ fractional part)

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$2^{\{\log_2 x\}}$ is not slowly varying ($\{\cdot\}$ fractional part) $\Rightarrow$ there is no limit theorem for $\frac{S_n - c_n}{a_n}$ for any choice of a_n, c_n .

There is on subsequences!

Theorem (Martin-Löf (1985))

$$\frac{S_{2^n}}{2^n} - n \xrightarrow{\mathcal{D}} W, \quad \text{as } n \rightarrow \infty.$$

W semistable rv. Moreover, convergence holds on subsequences $n_k = \lfloor \gamma 2^k \rfloor$, $\gamma \in (1/2, 1]$.

Theorem (Csörgő & Dodunekova (1991))

$\frac{S_{n_k}}{n_k} - \log_2 n_k$ converges in distribution if and only if

$$\gamma_{n_k} = \frac{n_k}{2^{\lceil \log_2 n_k \rceil}} \longrightarrow \gamma \in (1/2, 1].$$

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Merging

Theorem (Csörgő (2002))

$$\sup_{x \in \mathbb{R}} \left| \mathbf{P} \left\{ \frac{S_n}{n} - \log_2 n \leq x \right\} - G_{\gamma_n}(x) \right| \rightarrow 0, \quad \text{as } n \rightarrow \infty$$

The limit

Characteristic function of W_γ , $\gamma \in (1/2, 1]$,

$$\mathbf{E} \left(e^{itW_\gamma} \right) = \exp \left(ita + \int_0^\infty \left(e^{itx} - 1 - \frac{itx}{1+x^2} \right) dR_\gamma(x) \right),$$

with right-hand-side Lévy function

$$R_\gamma(x) = -\frac{\gamma}{2^{\lfloor \log_2(\gamma x) \rfloor}} = -\frac{2^{\{\log_2(\gamma x)\}}}{x}, \quad x > 0.$$

(semistable laws)

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Trimmed LLN

$X_1, X_2, \dots$ iid St. Petersburg rv's,

$$S_n = X_1 + \dots + X_n \quad \text{and} \quad X_n^* = \max_{1 \leq i \leq n} X_i$$

Theorem (Csörgő and Simons (1996))

$$\lim_{n \rightarrow \infty} \frac{S_n - X_n^*}{n \log_2 n} = 1 \quad a.s.$$

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Merging again

For $\gamma \in (1/2, 1]$ ($\approx$ Fréchet)

$$H_\gamma(x) = \begin{cases} 0, & \text{for } x \leq 0, \\ \exp(-\gamma 2^{-\lfloor \log_2(\gamma x) \rfloor}), & \text{for } x > 0. \end{cases}$$

Theorem (Berkes, Csáki & Csörgő (1999))

$$\sup_{x \in \mathbb{R}} \left| \mathbf{P} \left\{ \frac{X_n^*}{n} \leq x \right\} - H_{\gamma_n}(x) \right| = O(n^{-1}), \quad \text{as } n \rightarrow \infty.$$

Typical value: $X_n^* \approx 2^{\lceil \log_2 n \rceil + j}, j \in \mathbb{Z}.$

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Maximum and sum

Theorem (Darling (1952), Breiman (1965))

$Y, Y_1, Y_2, \dots$ iid ≥ 0 .

$$\frac{\max_{i \leq n} Y_i}{\sum_{i=1}^n Y_i} \xrightarrow{\mathcal{D}} Z,$$

with Z nondegenerate, iff $Y \in D(\alpha)$, $\alpha \in (0, 1)$; $Z = 1$ iff $\mathbf{P}\{Y > y\}$ is slowly varying, and $Z = 0$ iff $\sqrt{Y} \in D(2)$.

St. Petersburg case:

$$\frac{X_n^*}{S_n} \xrightarrow{\mathbf{P}} 0.$$

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For $j \in \mathbb{Z}$ and $\gamma \in [1/2, 1]$ introduce the notation

$$p_{j,\gamma} = e^{-\gamma 2^{-j}} \left(1 - e^{-\gamma 2^{-j}}\right), \quad \gamma_n = \frac{n}{2^{\lceil \log_2 n \rceil}}$$

Lemma

$$\sup_{j \in \mathbb{Z}} \left| \mathbf{P} \left\{ X_n^* = 2^{\lceil \log_2 n \rceil + j} \right\} - p_{j,\gamma_n} \right| = O(n^{-1}).$$

In particular for any $j \in \mathbb{Z}$, as $n \rightarrow \infty$

$$\mathbf{P} \left\{ X_n^* = 2^{\lceil \log_2 n \rceil + j} \right\} \sim e^{-\gamma_n 2^{-j}} \left(1 - e^{-\gamma_n 2^{-j}}\right).$$

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Small maximum

Put $N_n = |\{k : 1 \leq k \leq n, X_k = X_n^*\}|$.

Proposition

Conditionally on $X_n^ = 2^{k_n}$, where $\log_2 n - k_n \rightarrow \infty$*

$$\frac{N_n - \mathbf{E}[N_n | X_n^* = 2^{k_n}]}{\sqrt{\mathbf{Var}(N_n | X_n^* = 2^{k_n})}} \xrightarrow{\mathcal{D}} N(0, 1), \quad \text{as } n \rightarrow \infty.$$

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Typical maximum

Proposition (Gut & Martin-Löf (2016))

Conditionally on $X_n^* = 2^{\lceil \log_2 n \rceil + j}$, $j \in \mathbb{Z}$,

$$N_n \xrightarrow{\mathcal{D}} M_{j, \gamma n} \quad (\text{in the merging sense}),$$

where $M_{j, \gamma} \sim \text{Poisson}(2^{-j} \gamma)$ conditioned on not being zero.

Large maximum

Proposition

While, if $k_n - \log_2 n \rightarrow \infty$ then conditionally on $X_n^* = 2^{k_n}$

$$N_n \xrightarrow{\mathbf{P}} 1, \quad \text{as } n \rightarrow \infty.$$

Notation

$$\mathbf{P}\{X = 2^i | X \leq 2^k\} = 2^{-i} / (1 - 2^{-k})$$

$$F_k(x) = \mathbf{P}\{X \leq x | X \leq 2^k\} = \begin{cases} \frac{1}{1-2^{-k}} \left[1 - \frac{2^{\{\log_2 x\}}}{x} \right], & x \in [2, 2^k], \\ 1, & x > 2^k. \end{cases}$$

$X^{(k)}, X_1^{(k)}, \dots$, are iid F_k , and

$$S_n^{(k)} = X_1^{(k)} + \dots + X_n^{(k)}$$

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Conditioning on small maximum

Proposition

Given that $X_n^* = 2^{k_n}$, $k_n \geq 2$, such that $\log_2 n - k_n \rightarrow \infty$

$$\frac{S_n - \mathbf{E}[S_n | X_n^* = 2^{k_n}]}{\sqrt{\mathbf{Var}(S_n | X_n^* = 2^{k_n})}} \xrightarrow{\mathcal{D}} N(0, 1).$$

Conditioning on typical maximum

Proposition

$$\frac{S_{n_k}^{(\lceil \log_2 n_k \rceil + j)}}{n_k} - \log_2 n_k$$

converges in distribution iff $\gamma_{n_k} \rightarrow \gamma$. The limit $W_{j,\gamma}$

$$\varphi_{j,\gamma}(t) = \mathbf{E} e^{itW_{j,\gamma}} = \exp \left[itu_{j,\gamma} + \int_0^\infty (e^{itx} - 1 - itx) dL_{j,\gamma}(x) \right],$$

with

$$L_{j,\gamma}(x) = \begin{cases} \gamma 2^{-j} - \frac{2^{\{\log_2(\gamma x)\}}}{x}, & \text{for } x < 2^j \gamma^{-1}, \\ 0, & \text{for } x \geq 2^j \gamma^{-1}, \end{cases}$$

Conditioning on typical maximum

Proposition

For $j \in \mathbb{Z}$ we have

$$\left| \mathbf{P} \left\{ \frac{S_n}{n} - \log_2 n \leq x \mid X_n^* = 2^{\lceil \log_2 n \rceil + j} \right\} - \tilde{G}_{j, \gamma_n}(x) \right| \rightarrow 0,$$

where

$$\tilde{G}_{j, \gamma}(x) = \sum_{m=1}^{\infty} G_{j-1, \gamma} \left(x - m \frac{2^j}{\gamma} \right) \frac{(2^{-j} \gamma)^m}{m!} (e^{2^{-j} \gamma} - 1)^{-1}.$$

Corollary

Theorem (Gut & Martin-Löf (2016))

For any $\gamma \in [1/2, 1]$

$$G_\gamma(x) = \sum_{j=-\infty}^{\infty} \tilde{G}_{j,\gamma}(x) e^{-\gamma 2^{-j}} \left(1 - e^{-\gamma 2^{-j}}\right).$$

This is equivalent to the distributional representation

$$W_\gamma \stackrel{\mathcal{D}}{=} W_{Y_\gamma-1,\gamma} + M_{Y_\gamma,\gamma} 2^{Y_\gamma} \gamma^{-1},$$

where $(W_{j,\gamma})_{j \in \mathbb{Z}}$, $(M_{j,\gamma})_{j \in \mathbb{Z}}$ and Y_γ are independent,
 $Y_\gamma \sim (p_{j,\gamma})_{j \in \mathbb{Z}}$, $M_{j,\gamma} \sim \text{Poisson}(\gamma 2^{-j})$ conditioned on not being 0.

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Buchmann, Fan & Maller (2016) result

Lévy process setup: W_γ is a semistable Lévy process at time 1.

$$W_\gamma \stackrel{\mathcal{D}}{=} W_{Y_\gamma-1,\gamma} + M_{Y_\gamma,\gamma} 2^{Y_\gamma} \gamma^{-1},$$

The value $2^{Y_\gamma}/\gamma$ corresponds to the maximum jump, $M_{Y_\gamma,\gamma}$ is the number of the maximum jumps, and $W_{Y_\gamma-1,\gamma}$ has the law of the Lévy process conditioned on that the maximum jump is strictly less than $2^{Y_\gamma}/\gamma$.

This kind of distributional representations for general Lévy processes were obtained by Buchmann, Fan & Maller (2016).

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Conditioning on large maximum

Proposition

Assume that $k_n - \log_2 n \rightarrow \infty$. Given that $X_n^* = 2^{k_n}$

$$\frac{S_n}{X_n^*} - A_n \xrightarrow{\mathbf{P}} 1,$$

where

$$A_n = \frac{nk_n}{2^{k_n}}.$$

Conditioning on the maximum

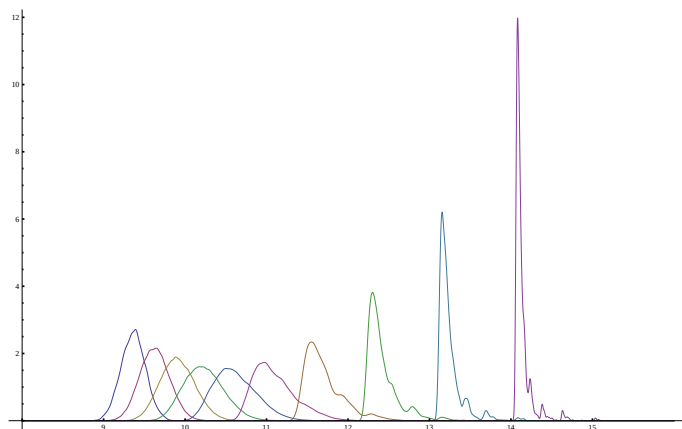


Figure: The conditional histograms for $\log_2 S_n$, $n = 2^7$

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Joint with István Berkes and László Györfi.

Subexponential distributions

$Y, Y_1, Y_2, \dots$ iid ≥ 0 , $G, \overline{G}(x) = 1 - G(x)$.

G is *subexponential*, $G \in \mathcal{S}$,

$$\lim_{x \rightarrow \infty} \frac{\overline{G * G}(x)}{\overline{G}(x)} = 2,$$

Characterizing property of $\mathcal{S}$: for any $n \geq 1$

$$\lim_{x \rightarrow \infty} \frac{\mathbf{P}\{Y_1 + \dots + Y_n > x\}}{\mathbf{P}\{\max\{Y_i : i = 1, 2, \dots, n\} > x\}} = 1,$$

equivalently

$$\lim_{x \rightarrow \infty} \frac{\mathbf{P}\{Y_1 + \dots + Y_n > x\}}{\mathbf{P}\{Y_1 > x\}} = n.$$

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O-subexponential distributions

Goldie (1978): St. Petersburg distribution F is not subexponential.

$$2 = \liminf_{x \rightarrow \infty} \frac{\overline{F * F}(x)}{\overline{F}(x)} < \limsup_{x \rightarrow \infty} \frac{\overline{F * F}(x)}{\overline{F}(x)} = 4.$$

G is *O-subexponential* (Klüppelberg, 1990), $G \in \mathcal{OS}$, if

$$\ell^*(G) := \limsup_{x \rightarrow \infty} \frac{\overline{G * G}(x)}{\overline{G}(x)} < \infty.$$

always $\liminf \geq 2$; $= 2$ for heavy-tailed (Foss & Korshunov 2007)

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G is *O-subexponential* (Klüppelberg, 1990), $G \in \mathcal{OS}$, if

$$\ell^*(G) := \limsup_{x \rightarrow \infty} \frac{\overline{G * G}(x)}{\overline{G}(x)} < \infty.$$

always $\liminf \geq 2$; $= 2$ for heavy-tailed (Foss & Korshunov 2007)

Shimura and Watanabe (2005): $G \in \mathcal{OS}$, $\forall \varepsilon > 0$, $\exists c > 0$,

$$\frac{\overline{G^{n*}}(x)}{\overline{G}(x)} \leq c(\ell^*(G) - 1 + \varepsilon)^n.$$

Notation

$X, X_1, X_2, \dots$ iid St. Petersburg rv's

$X_{1n} \leq X_{2n} \leq \dots \leq X_{nn}$ ordered sample of $X_1, X_2, \dots, X_n$.

r -trimmed sum: $S_{n,r} = \sum_{k=1}^{n-r} X_{kn}$.

Tail of the sums

Theorem

As n, r fix, $x \rightarrow \infty$

$$\mathbf{P}\{S_{n,r} > x\} \sim \frac{2^{(r+1)\{\log_2 x\}}}{x^{r+1}} \binom{n}{r+1} \times \left(1 + \mathbf{P}\left\{S_{n-r-1} > x(1 - 2^{-\{\log_2 x\}})\right\} (2^{r+1} - 1)\right).$$

In particular, for any $0 < \delta < 1$,

$$\lim_{x \rightarrow \infty, \{\log_2 x\} > \delta} \mathbf{P}\{S_{n,r} > x\} \frac{x^{r+1}}{2^{(r+1)\{\log_2 x\}}} = \binom{n}{r+1}.$$

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$$r = 0$$

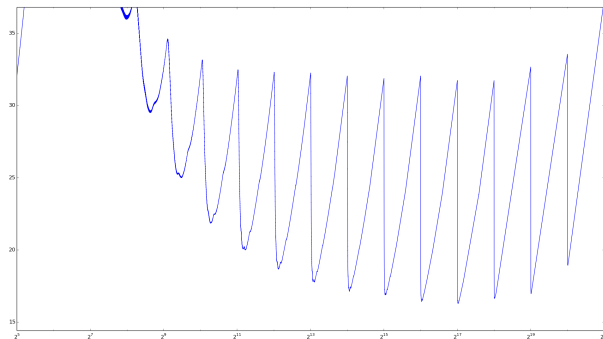


Figure: The function $x \cdot \mathbf{P}\{S_{16} > x\}$ in a logarithmic scale.

Almost subexponential

Untrimmed case:

$$\mathbf{P}\{S_n > x\} \sim \frac{2^{\{\log_2 x\}}}{x} n \left(1 + \mathbf{P}\{S_{n-1} > x(1 - 2^{-\{\log_2 x\}})\} \right)$$

from which

$$n = \liminf_{x \rightarrow \infty} x \mathbf{P}\{S_n > x\} < \limsup_{x \rightarrow \infty} x \mathbf{P}\{S_n > x\} = 2n.$$

Since $x \mathbf{P}\{X > x\} = 2^{\{\log_2 x\}}$, $x \geq 2$, we have

$$\lim_{x \rightarrow \infty, \{\log_2 x\} \geq \delta} \frac{\mathbf{P}\{S_n > x\}}{\mathbf{P}\{X > x\}} = n.$$

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Outline

St. Petersburg game

Sum

Maximum

Conditioning on the maximum

Number of maximum terms

Conditional limit results

Trimmed sums

Finite number of summands

Properties of the r -trimmed limit

$E_k, k = 1, 2, \dots$ iid $\text{Exp}(1)$, $Z_k = E_1 + \dots + E_k$

Theorem

Let $n_k = \lfloor \gamma 2^k \rfloor$, for some $\gamma \in (1/2, 1]$. Then for any $r \geq 0$

$$\frac{1}{n_k} S_{n_k, r} - a_{n_k, \gamma}^{(r)} \xrightarrow{\mathcal{D}} Y_{r, \gamma} = \sum_{k=r+1}^{\infty} \gamma^{-1} \left(2^{-\lfloor \log_2 Z_k / \gamma \rfloor} - 2^{-\lfloor \log_2 k / \gamma \rfloor} \right),$$

with centering sequence

$$a_{n, \gamma}^{(r)} = \gamma^{-1} \sum_{j=r+1}^n 2^{-\lfloor j / \gamma \rfloor}.$$

Proof (sketch)

Quantile method & LePage, Woodroffe, Zinn idea.

Quantile representation: $(X_{1n}, \dots, X_{nn}) \stackrel{\mathcal{D}}{=} (Q(U_{1n}), \dots, Q(U_{nn}))$,
where $F^{-1}(s) = Q(s) = \inf\{x : s \leq F(x)\}$

$$Q(s) = \begin{cases} 2, & s = 0, \\ 2^{\lceil -\log_2(1-s) \rceil} = \frac{2^{\{\log_2(1-s)\}}}{1-s}, & s \in (0, 1). \end{cases}$$

$(E_i)_{i \in \mathbb{N}}$ iid $\text{Exp}(1)$, $Z_n = E_1 + \dots + E_n$. For n fix

$$(U_{1n}, U_{2n}, \dots, U_{nn}) \stackrel{\mathcal{D}}{=} \left(\frac{Z_1}{Z_{n+1}}, \frac{Z_2}{Z_{n+1}}, \dots, \frac{Z_n}{Z_{n+1}} \right),$$

where U 's are ordered sample of n iid $\text{Uniform}(0, 1)$.

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Proof

$\Psi(x) = 2^{\{\log_2 x\}}$ (grows linearly from 1 to 2 in each $[2^j, 2^{j+1})$).

$$Q(1-s) = \Psi(s)/s$$

$$(X_{1n}, \dots, X_{nn}) \stackrel{\mathcal{D}}{=} \left(\frac{Z_{n+1}}{Z_1} \Psi(Z_1/Z_{n+1}), \dots, \frac{Z_{n+1}}{Z_n} \Psi(Z_n/Z_{n+1}) \right)$$

SLLN $Z_{n+1}/n \rightarrow 1$ a.s.

$$X_{j,n}^* = \frac{n}{Z_j} \Psi\left(\frac{Z_j}{n}\right) (1 + o(1)) \quad \text{a.s.}$$

$$\Psi(Z_j/n) = \Psi(Z_j/\gamma_n)$$

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LePage, Woodroffe & Zinn (1981)

$Y, Y_1, Y_2, \dots$ iid, ≥ 0 , $Y \in D(\alpha)$, S_n partial sum,
 $(S_n - nb_n)/a_n \rightarrow S$. $Y_{1,n} \geq Y_{2,n} \geq \dots \geq Y_{n,n}$

$$S = \sum_{k=1}^{\infty} \left(Z_k^{-1/\alpha} - \mathbf{E} Z_k^{-1/\alpha} I(Z_k^{-1/\alpha} < 1) \right),$$

where $E_1, E_2, \dots$ are iid $\text{Exp}(1)$, $Z_k = E_1 + \dots + E_k$. Moreover,

$$\left(\frac{S_n - nb_n}{a_n}, \left(\frac{Y_{1,n}}{a_n}, \dots, \frac{Y_{n,n}}{a_n} \right) \right) \xrightarrow{\mathcal{D}} \left(S, (Z_1^{-1/\alpha}, Z_2^{-1/\alpha}, \dots) \right).$$

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On the centering

For any $\gamma \in (1/2, 1]$, $n_k = \lfloor \gamma 2^k \rfloor$,

$$a_{n_k, \gamma}^{(0)} - \log_2 n_k \rightarrow 2 - \frac{1}{\gamma} \sum_{k=1}^{\infty} \frac{k \varepsilon_k}{2^k} - \log_2 \gamma = \xi(\gamma),$$

where $\gamma = \sum_{k=1}^{\infty} \varepsilon_k 2^{-k}$.

Steinhaus' resolution of the St. Petersburg paradox (Csörgő & Simons 1993)

ξ is right-continuous, left-continuous except at dyadic rationals greater than $1/2$ and has unbounded variation (Csörgő & Simons 1993); the Hausdorff and box-dimension of the graph of ξ is 1 (Kern & Wedrich 2014).

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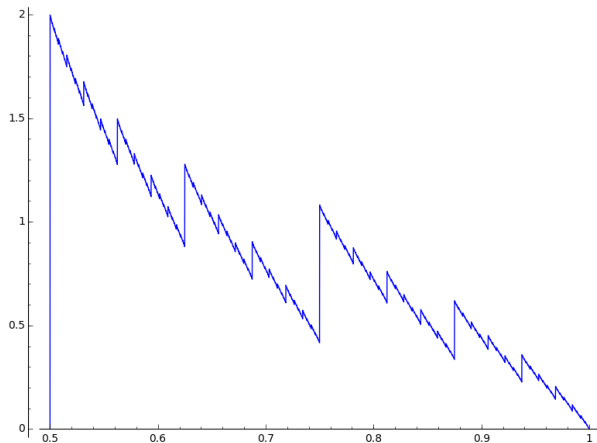
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Properties of the r -trimmed limit

$$\xi(\gamma)$$



Tail of the trimmed limit

$$A_{r,\gamma} = \gamma^{-1} \sum_{k=1}^r 2^{\lfloor k/\gamma \rfloor}$$

Theorem

$$\begin{aligned} \mathbf{P}\{Y_{r,\gamma} > x\} &\sim \frac{2^{\{\log_2(\gamma x)\}}(r+1)}{(r+1)! x^{r+1}} \left[2^{-r-1} + (2^{r+1} - 1) \right. \\ &\quad \times \left. \sum_{\ell=0}^1 2^{-\ell(r+1)} \mathbf{P}\left\{Y_{0,\gamma} + A_{r,\gamma} > x \left(1 - 2^{\ell - \{\log_2(\gamma x)\}}\right)\right\} \right]. \end{aligned}$$

Untrimmed case

$$\mathbf{P}\{Y_{0,\gamma} > x\} \sim \frac{2^{\{\log_2(\gamma x)\}}}{x} \times \left[2^{-1} + \sum_{\ell=0}^1 2^{-\ell} \mathbf{P}\left\{Y_{0,\gamma} > x \left(1 - 2^{\ell - \{\log_2(\gamma x)\}}\right)\right\} \right].$$

Exactly the tail of the Lévy measure appears

$$\frac{\mathbf{P}\{Y_{0,\gamma} > x\}}{-R_\gamma(x)} \sim 2^{-1} + \sum_{\ell=0}^1 2^{-\ell} \mathbf{P}\left\{Y_{0,\gamma} > x \left(1 - 2^{\ell - \{\log_2(\gamma x)\}}\right)\right\} \right].$$

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For any $\delta \in (0, 1/2)$ we have

$$\lim_{x \rightarrow \infty, \delta < \{\log_2(\gamma x)\} < 1 - \delta} \mathbf{P}\{Y_{r,\gamma} > x\} \frac{x}{2^{\{\log_2(\gamma x)\}}} = 1.$$

In the untrimmed case ($r = 0$) for $\gamma = 1$

$$\mathbf{P}\{Y_{0,1} > 2^m + c\} \sim 2^{-m} [1 + \mathbf{P}\{Y_{0,1} > c\}], \quad \text{as } m \rightarrow \infty,$$

(Martin-Löf 1985).

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Watanabe & Yamamuro (2012) result

For general semistable distributions:

$$\lim_{n \rightarrow \infty} 2^n \mathbf{P}\{W_1 > x 2^n\} = -R_1(x) + [R_1(x-) - R_1(x)] \mathbf{P}\{W_1 > 0\}$$

$$C_* = \liminf_{x \rightarrow \infty} \frac{\mathbf{P}\{W > x\}}{-R(x)} \leq \limsup_{x \rightarrow \infty} \frac{\mathbf{P}\{W > x\}}{-R(x)} = C^*,$$

with

$$C_* = 1 - (1 - Q^{-1}) \mathbf{P}\{W < 0\}, \quad C^* = Q + (Q - 1) \mathbf{P}\{W < 0\},$$

$$\text{and } Q = \sup_{x \in [1, 2]} R(x-)/R(x).$$

‘the modern student will hardly understand the mysterious discussions of this “paradox” ’ – Feller

A natural example, which is not in the domain of attraction of any stable law, but it is in the domain of geometric partial attraction of a semistable law.

Not subexponential, but tractable tail behavior.

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References I



I. Berkes, L. Györfi, P. Kevei

Tail probabilities of St. Petersburg sums, trimmed sums, and their limit. *Journal of Theoretical Probability*. To appear.



G. Fukker, L. Györfi, P. Kevei

Asymptotic behavior of the St. Petersburg sum conditioned on its maximum. *Bernoulli*, **22** (2), 1026-1054, 2016.